

# Large Deviations and Importance Sampling for Weakly Interacting Diffusions\*

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# Overview

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Importance  
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Small-Noise  
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## 2 Importance Sampling

- Small-Noise Diffusions
- Weakly Interacting Diffusions

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# Setup and Motivation: Weakly Interacting Diffusions

We consider particles  $\{X_t^{i,N}\}_{i=1}^N$  obeying the system of  $N$  SDEs:

$$dX_t^{i,N} = b(X_t^{i,N}, \mu_t^N)dt + \sigma(X_t^{i,N}, \mu_t^N)dW_t^i, \quad X_0^{i,N} = x^{i,N},$$

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where  $\{W^i\}_{i=1}^N$  are independent  $m$ -dimensional standard Brownian motions,

$b : \mathbb{R}^d \times \mathcal{P}(\mathbb{R}^d) \rightarrow \mathbb{R}^d$ ,  $\sigma : \mathbb{R}^d \times \mathcal{P}(\mathbb{R}^d) \rightarrow \mathbb{R}^{d \times m}$ , and

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$$\mu_t^N := \frac{1}{N} \sum_{i=1}^N \delta_{X_t^{i,N}}, \quad t \in [0, T], \quad \text{"Empirical Measure".}$$

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## Applications:

- mathematical biology and neuroscience (Delarue et al., 2015)
- machine learning (V. Nilsson and P. Nyquist, 2022)
- stochastic portfolio theory (E. R. Fernholz, 2002)
- systemic risk/clustering of defaults (J. Garnier et al., 2013; P. Dai Pra et al., 2009)
- many more (Chaintron and Diez, 2022)

# Propagation of Chaos

Theorem: Propagation of Chaos (M. Kac, 1957; H.P. McKean, 1969)

Let

$$dX_t^{i,N} = b(X_t^{i,N}, \mu_t^N)dt + \sigma(X_t^{i,N}, \mu_t^N)dW_t^i, \quad X_0^{i,N} = x^{i,N} \quad (1)$$

$$\mu_t^N := \frac{1}{N} \sum_{i=1}^N \delta_{X_t^{i,N}}.$$

Then  $\mu^N$  converges weakly as  $N \rightarrow \infty$  as a  $C([0, T]; \mathcal{P}(\mathbb{R}^d))$ -valued random variable to the deterministic process  $\mathcal{L}(X)$ , where  $X_t$  satisfies:

$$dX_t = b(X_t, \mathcal{L}(X_t))dt + \sigma(X_t, \mathcal{L}(X_t))dW_t, \quad X_0 \sim \nu \quad (2)$$

$$\nu := \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{i=1}^N \delta_{X^{i,N}}.$$

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Further suppose  $x^{i,N}$  are IID with law  $\nu$ . Then as  $N \rightarrow \infty$ , any  $k$  “tagged” particles obeying (1) converge to independent processes obeying Equation (2).

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# The Classical Model of Dawson

(D.A. Dawson, 1983) considered the one-dimensional system:

$$dX_t^{i,N} = \left[ -(X_t^{i,N})^3 + X_t^{i,N} - \theta \left( X_t^{i,N} - \frac{1}{N} \sum_{j=1}^N X_t^{j,N} \right) \right] dt + \sigma dW_t^i$$

$i \in \{1, \dots, N\}$ ,  $\theta, \sigma > 0$ . This corresponds to  
 $b(x, \mu) = -x^3 + x - \theta(x - \int z \mu(dz))$ .

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 $b(x, \mu) = -x^3 + x - \theta(x - \int z \mu(dz))$ . It was shown:

- The system undergoes the propagation of chaos
- There exists a critical temperature  $\sigma_c > 0$  such that:

Temperature ( $\sigma$ )	Finite $N$ Particles	$N \rightarrow \infty$ Limit
$\sigma \geq \sigma_c$	$\exists!$ invariant density	$\exists!$ invariant density
$0 < \sigma < \sigma_c$	$\exists!$ invariant density	3 invariant densities

# The Classical Model of Dawson

$$dX_t = \left[ -(X_t)^3 + X_t - \theta(X_t - \mathbb{E}[X_t]) \right] dt + \sigma dW_t$$

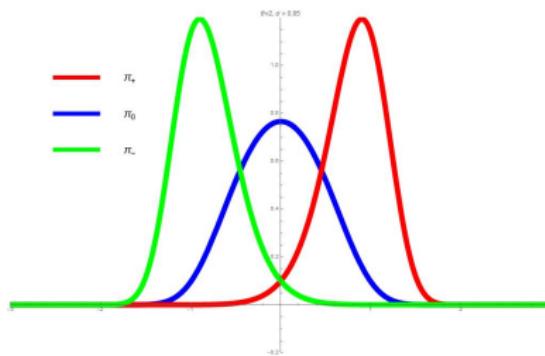
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Suppose  $\mu_0^N \rightarrow \pi_+$ . Then for  $T > 0$  and  $\pi_+ \notin A \subset \mathcal{P}(\mathbb{R}^d)$ :

$$\mathbb{P}(\mu_T^N \in A) \rightarrow 0 \text{ as } N \rightarrow \infty.$$

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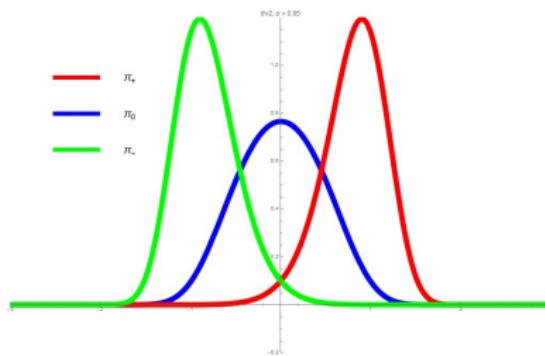
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For example:

$$\mathbb{P}\left(\frac{1}{N} \sum_{i=1}^N X_T^{i,N} < 0\right) \rightarrow 0 \text{ as } N \rightarrow \infty.$$

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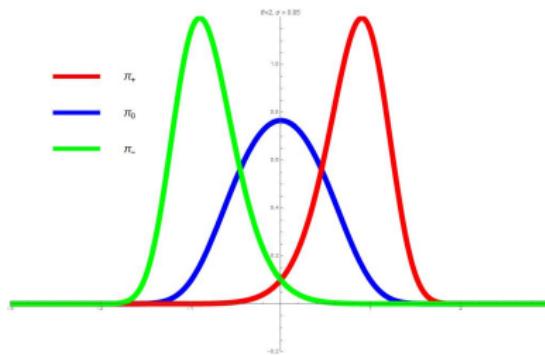
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For example:

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Our long-term goal: Efficiently estimate such a quantity for large  $N^{6/23}$

# The Problem with Standard Monte Carlo (1)

- We will simplify the problem to estimating  $\mathbb{E}[\exp(-NG(X^N))]$  for  $G \in C_b(\mathcal{X})$  and  $\{X^N\} \subset \mathcal{X}$ .

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- Let  $\hat{X}^N$  have Law  $\hat{\mathbb{P}}^N$  and  $X^N$  have Law  $\mathbb{P}^N$ . Consider the unbiased estimator:

$$\hat{\delta}(N) = \frac{1}{M} \sum_{j=1}^M \exp(-NG(\hat{X}^{N,j})) \frac{d\mathbb{P}^N}{d\hat{\mathbb{P}}^N}(\hat{X}^{N,j}),$$

where  $\{\hat{X}^{N,j}\}_{j=1}^M$  are  $M$  independent samples of  $\hat{X}^N$ .

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where  $\{\hat{X}^{N,j}\}_{j=1}^M$  are  $M$  independent samples of  $\hat{X}^N$ .

- Measure performance via relative error:

$$\rho(\hat{\delta}(N)) = \frac{\sqrt{\text{Var}(\hat{\delta}(N))}}{\mathbb{E}[\hat{\delta}(N)]} = \frac{1}{\sqrt{M}} \sqrt{R(\hat{\delta}(N)) - 1}$$

$$R(\hat{\delta}(N)) = \mathbb{E} \left[ \exp(-2NG(\hat{X}^N)) \left( \frac{d\mathbb{P}^N}{d\hat{\mathbb{P}}^N}(\hat{X}^N) \right)^2 \right] / \mathbb{E}[\exp(-NG(X^N))]^2$$

# The Problem with Standard Monte Carlo (2)

## Definition

We say that a sequence of  $\mathcal{X}$ -valued random variables  $\{X^N\}_{N \in \mathbb{N}}$  satisfies the *Laplace Principle* (equivalently LDP) with rate function  $I : \mathcal{X} \rightarrow [0, +\infty]$  if for all  $G \in C_b(\mathcal{X})$ ,

$$\lim_{N \rightarrow \infty} \frac{1}{N} \log \mathbb{E} \left[ \exp \left( -NG(X^N) \right) \right] = - \inf_{x \in \mathcal{X}} \left[ G(x) + I(x) \right].$$

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Suppose  $\{X^N\}$  satisfies the LDP with rate function  $I$ . Then:

$$\lim_{N \rightarrow \infty} \frac{1}{N} \log \mathbb{E} [\exp(-2NG(X^N))] = - \inf_{x \in \mathcal{X}} \left[ 2G(x) + I(x) \right] := -\gamma_2,$$

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Letting  $\delta(N)$  be the standard M.C. Estimator ( $\hat{X}^N = X^N$ ):

$$\rho(\delta(N)) = \frac{1}{\sqrt{M}} \sqrt{R(\delta(N)) - 1}, \quad R(\delta(N)) = \frac{\mathbb{E}[\exp(-2NG(X^N))]}{\mathbb{E}[\exp(-NG(X^N))]^2}$$

$$\lim_{N \rightarrow \infty} \frac{1}{N} \log R(\delta(N)) = 2\gamma_1 - \gamma_2 \geq 0$$

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$$\lim_{N \rightarrow \infty} \frac{1}{N} \log R(\delta(N)) = 2\gamma_1 - \gamma_2 \geq 0$$

- So Standard M.C. has exponentially poor performance in  $N$ .

# Log Efficiency

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Definition: Log-Efficient

$\hat{\delta}(N)$  is called log-efficient if

$$\lim_{N \rightarrow \infty} -\frac{1}{N} \log R(\hat{\delta}(N)) = 0.$$

# Example: Small-Noise Diffusions (1)

- Let  $\mathcal{X} = C([0, T]; \mathbb{R}^d)$ ,  $G(\phi) = g(\phi(T))$ , and  $X^N$  be given by

$$dX_t^N = b(X_t^N)dt + \frac{1}{\sqrt{N}}\sigma(X_t^N)dW_t.$$

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- A natural change of measure is given by Girsanov's Theorem, so

$$d\hat{X}_t^{N,v} = [b(\hat{X}_t^{N,v}) + \sigma(\hat{X}_t^{N,v})v(t, \hat{X}_t^{N,v})]dt + \frac{1}{\sqrt{N}}\sigma(\hat{X}_t^{N,v})dW_t$$

$$\frac{d\mathbb{P}^N}{d\hat{\mathbb{P}}^{N,v}}(\hat{X}^{N,v}) = \exp\left(-\sqrt{N}\int_0^T v(t, \hat{X}_t^{N,v}) \cdot dW_t - \frac{N}{2} \int_0^T |v(t, \hat{X}_t^{N,v})|^2 dt\right)$$

for some  $v : [0, T] \times \mathbb{R}^d \rightarrow \mathbb{R}^m$ .

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for some  $v : [0, T] \times \mathbb{R}^d \rightarrow \mathbb{R}^m$ .

- Letting  $v^N = -\sigma^\top \nabla_x \Phi^N$ , where:

$$-\partial_t \Phi^N - b \cdot \nabla_x \Phi^N + \frac{1}{2}|\sigma^\top \nabla_x \Phi^N|^2 - \frac{1}{2N}\sigma\sigma^\top : \nabla_x^2 \Phi^N = 0$$

$$\Phi^N(T, x) = g(x),$$

we have  $v^N$  provides a zero-variance estimator.

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- Sending  $N \rightarrow \infty$ , we expect that  $v = -\sigma^\top \nabla_x \Psi$  for  $\Psi$  solving the associated zero-viscosity HJB Equation will yield small relative error uniformly in  $N$ .

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# Example: Small-Noise Diffusions (2)

$$dX_t^N = b(X_t^N)dt + \frac{1}{\sqrt{N}}\sigma(X_t^N)dW_t$$

$$d\hat{X}_t^{N,v} = [b(\hat{X}_t^{N,v}) + \sigma(\hat{X}_t^{N,v})v(t, \hat{X}_t^{N,v})]dt + \frac{1}{\sqrt{N}}\sigma(\hat{X}_t^{N,v})dW_t$$

$$-\partial_t \Psi - b \cdot \nabla_x \Psi + \frac{1}{2} |\sigma^\top \nabla_x \Psi|^2 = 0, \quad \Psi(T, x) = g(x).$$

Theorem (E. Vanden-Eijnden and J. Weare, 2012)

The importance sampling scheme for estimating  $\mathbb{E}[\exp(-Ng(X_T^N))]$  using  $v(t, x) = -\sigma^\top(x)\nabla_x \Psi(t, x)$  is log-efficient.

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$$\Psi(s, x) = \inf_{\phi \in C([s, T]; \mathbb{R}^d) : \phi(s) = x} \left\{ J^s(\phi) + g(\phi(T)) \right\} \quad (3)$$

where  $J^s : C([s, T]; \mathbb{R}^d) \rightarrow [0, +\infty]$  is the LDP rate function for  $\{X^N\}$  starting at position  $x$  at time  $s$ :

$$\begin{aligned} J^s(\phi) &= \frac{1}{2} \int_s^T |(\dot{\phi}(t) - b(\phi(t)))^\top [\sigma\sigma^\top(\phi(t))]^{-1}(\dot{\phi}(t) - b(\phi(t)))| dt \\ &= \inf_{u: [s, T] \rightarrow \mathbb{R}^m: \dot{\phi}(t) = b(\phi(t)) + \sigma(\phi(t))u(t)} \frac{1}{2} \int_s^T |u(t)|^2 dt. \end{aligned}$$

# Back to Interacting Diffusions (1)

$$dX_t^{i,N} = b(X_t^{i,N}, \mu_t^N)dt + \sigma(X_t^{i,N}, \mu_t^N)dW_t^i, \quad \mu_t^N := \frac{1}{N} \sum_{i=1}^N \delta_{X_t^{i,N}}$$

- Want to estimate  $\mathbb{E}[\exp(-Ng(\mu_T^N))]$  for  $g : \mathcal{P}(\mathbb{R}^d) \rightarrow \mathbb{R}$ .
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Use again Girsanov:

$$\begin{aligned} d\hat{X}_t^{i,N,v} &= [b(\hat{X}_t^{i,N,v}, \hat{\mu}_t^{N,v}) + \sigma(\hat{X}_t^{i,N,v}, \hat{\mu}_t^{N,v})v_i(t, \hat{X}_t^{1,N,v}, \dots, \hat{X}_t^{N,N,v})]dt \\ &\quad + \sigma(\hat{X}_t^{i,N,v}, \hat{\mu}_t^{N,v})dW_t^i, \quad \hat{\mu}_t^{N,v} := \frac{1}{N} \sum_{i=1}^N \delta_{\hat{X}_t^{i,N,v}} \end{aligned}$$

$$\begin{aligned} \frac{\mathbb{P}^N}{\hat{\mathbb{P}}^{N,v}}(\hat{\mu}_T^{N,v}) &= \exp\left(-\sum_{i=1}^N \int_0^T v_i(t, \hat{X}_t^{1,N,v}, \dots, \hat{X}_t^{N,N,v}) \cdot dW_t^i\right. \\ &\quad \left.+ \frac{1}{2} \int_0^T |v_i(t, \hat{X}_t^{1,N,v}, \dots, \hat{X}_t^{N,N,v})|^2 dt\right) \end{aligned}$$

For  $v : [0, T] \times \mathbb{R}^{dN} \rightarrow \mathbb{R}^{mN}$ .

# Back to Interacting Diffusions (2)

$v_i^N(t, x_1, \dots, x_N) = -N\sigma^\top(x_i, \mu_x^N)\nabla_{x_i}\tilde{\Phi}^N(t, x_1, \dots, x_N)$  yields a zero-variance estimator, where  $\mu_x^N := \frac{1}{N}\sum_{i=1}^N \delta_{x_i}$  and:

$$\begin{aligned} -\partial_t \tilde{\Phi}^N(t, x) &= \sum_{i=1}^N \left\{ b(x_i, \mu_x^N) \cdot \nabla_{x_i} \tilde{\Phi}^N(t, x) - \frac{N}{2} |\sigma^\top(x_i, \mu_x^N)\nabla_{x_i}\tilde{\Phi}^N(t, x)|^2 \right. \\ &\quad \left. + \frac{1}{2}\sigma\sigma^\top(x_i, \mu_x^N) : \nabla_{x_i}^2 \tilde{\Phi}^N(t, x) \right\}, x \in \mathbb{R}^{Nd}, t \in [0, T] \end{aligned}$$

$$\tilde{\Phi}^N(T, x_1, \dots, x_N) = g(\mu_x^N).$$

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- Let  $\Phi^N : \mathcal{P}_2(\mathbb{R}^d) \rightarrow \mathbb{R}$  be such that  $\Phi^N(\mu_x^N) = \tilde{\Phi}^N(x_1, \dots, x_N)$ .

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- Question: How can we send  $N \rightarrow \infty$  to get a “zero-viscosity” HJB Equation?
- Let  $\Phi^N : \mathcal{P}_2(\mathbb{R}^d) \rightarrow \mathbb{R}$  be such that  $\Phi^N(\mu_x^N) = \tilde{\Phi}^N(x_1, \dots, x_N)$ .
- Then by Propositions 5.35 and 5.91 of (R. Carmona and F. Delarue, 2018):

$$\nabla_{x_i}\tilde{\Phi}^N(x) = \frac{1}{N}\partial_\mu\Phi^N(\mu_x^N)[x_i]$$

$$\nabla_{x_i}^2\tilde{\Phi}^N(x) = \frac{1}{N}\partial_z\partial_\mu\Phi^N(\mu_x^N)[x_i] + \frac{1}{N^2}\partial_\mu^2\Phi^N(\mu_x^N)[x_i, x_i]$$

- $\partial_\mu\Phi^N(\nu)[\cdot]$  denotes the **Lions derivative** (P. L. Lions 2012)

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# Back to Interacting Diffusions (3)

$$v_i^N(t, x_1, \dots, x_N) = -\sigma^\top(\mathbf{x}_i, \mu_x^N) \partial_\mu \Phi^N(t, \mu_x^N)[\mathbf{x}_i]$$

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# Back to Interacting Diffusions (3)

$$v_i^N(t, x_1, \dots, x_N) = -\sigma^\top(x_i, \mu_x^N) \partial_\mu \Phi^N(t, \mu_x^N)[x_i]$$

Plugging into the Equation  $\tilde{\Phi}^N$  solves,  $\Phi^N(T, \mu_x^N) = g(\mu_x^N)$ , and:

$$\begin{aligned} -\partial_t \Phi^N(t, \mu_x^N) &= \\ \frac{1}{N} \sum_{i=1}^N \left\{ b(x_i, \mu_x^N) \cdot \partial_\mu \Phi^N(t, \mu_x^N)[x_i] - \frac{1}{2} |\sigma^\top(x_i, \mu_x^N) \partial_\mu \Phi^N(t, \mu_x^N)[x_i]|^2 \right. \\ \left. + \frac{1}{2} \sigma \sigma^\top(x_i, \mu_x^N) : [\partial_z \partial_\mu \Phi^N(t, \mu_x^N)[x_i] + \frac{1}{N} \partial_\mu^2 \Phi^N(t, \mu_x^N)[x_i, x_i]] \right\}, x \in \mathbb{R}^{Nd}. \end{aligned}$$

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We rewrite the sum as an integral, so:

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Now we have all terms only depend on the measure  $\mu_x^N = \frac{1}{N} \sum_{i=1}^N \delta_{x_i}$ , and we can send  $N \rightarrow \infty$ !

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$$\begin{aligned}-\partial_t \Psi(t, \nu) = & \int_{\mathbb{R}^d} b(z, \nu) \cdot \partial_\mu \Psi(t, \nu)[z] - \frac{1}{2} |\sigma^\top(z, \nu) \partial_\mu \Psi(t, \nu)[z]|^2 \\ & + \frac{1}{2} \sigma \sigma^\top(z, \nu) : \partial_z \partial_\mu \Psi(t, \nu)[z] \nu(dz), \quad \Psi(T, \nu) = g(\nu), \nu \in \mathcal{P}_2(\mathbb{R}^d)\end{aligned}$$

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Theorem (Z. B. and M. Heldman, 2022)

The IS scheme for estimating  $\mathbb{E}[\exp(-Ng(\mu_T^N))]$  using  $v(t, \nu, z) = -\sigma^\top(z, \nu) \partial_\mu \Psi(t, \nu)[z]$  is log-efficient.

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$$\Psi(s, \nu) = \inf_{\mu \in C([s, T]; \mathcal{P}(\mathbb{R}^d)) : \mu(s) = \nu} \{I^s(\mu) + g(\mu_T)\}$$

where  $I^s : C([s, T]; \mathcal{P}(\mathbb{R}^d)) \rightarrow [0, +\infty]$  is the LDP rate function for  $\mu^N$  with initial distribution  $\mu_s^N \rightarrow \nu$ .

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# Main Theorem

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$$I^s(\mu) = \frac{1}{2} \int_s^T \sup_{\phi \in C_c^\infty(\mathbb{R}^d): \text{den} \neq 0} \frac{|\langle \dot{\mu}(t) - L_{\mu(t)}^* \mu(t), \phi \rangle|^2}{\langle \mu(t), \|\nabla^\top \phi \sigma \sigma^\top \nabla \phi\|^2 \rangle} dt$$

$$L_\nu \phi(x) = b(x, \nu) \cdot \nabla \phi(x) + \frac{1}{2} \sigma \sigma^\top(x, \nu) : \nabla^2 \phi(x)$$

$$I^s(\mu) = \inf_{u \in \mathcal{U}: \mu(t) = \mathcal{L}(\hat{X}_t^u), t \in [s, T]} \frac{1}{2} \mathbb{E} \left[ \int_s^T |u(t)|^2 dt \right]$$

$$d\hat{X}_t^u = [b(\hat{X}_t^u, \mathcal{L}(\hat{X}_t^u)) + \sigma(\hat{X}_t^u, \mathcal{L}(\hat{X}_t^u)) u(t)] dt + \sigma(\hat{X}_t^u, \mathcal{L}(\hat{X}_t^u)) dW_t, \quad \hat{X}_s^u \sim \nu$$

- Rate function representations from (D. A. Dawson and J. Gärtner, 1987) and (A. Budhiraja et al., 2012)

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$$I^s(\mu) = \frac{1}{2} \int_s^T \sup_{\phi \in C_c^\infty(\mathbb{R}^d): \text{den} \neq 0} \frac{|\langle \dot{\mu}(t) - L_{\mu(t)}^* \mu(t), \phi \rangle|^2}{\langle \mu(t), \|\nabla^\top \phi \sigma \sigma^\top \nabla \phi\|^2 \rangle} dt$$

$$L_\nu \phi(x) = b(x, \nu) \cdot \nabla \phi(x) + \frac{1}{2} \sigma \sigma^\top(x, \nu) : \nabla^2 \phi(x)$$

$$I^s(\mu) = \inf_{u \in \mathcal{U}: \mu(t) = \mathcal{L}(\hat{X}_t^u), t \in [s, T]} \frac{1}{2} \mathbb{E} \left[ \int_s^T |u(t)|^2 dt \right]$$

$$d\hat{X}_t^u = [b(\hat{X}_t^u, \mathcal{L}(\hat{X}_t^u)) + \sigma(\hat{X}_t^u, \mathcal{L}(\hat{X}_t^u)) u(t)] dt + \sigma(\hat{X}_t^u, \mathcal{L}(\hat{X}_t^u)) dW_t, \quad \hat{X}_s^u \sim \nu$$

- Rate function representations from (D. A. Dawson and J. Gärtner, 1987) and (A. Budhiraja et al., 2012)
- Assuming sufficient regularity for an expansion of the prelimit HJB equation, can further prove vanishing relative error in  $N$

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$$\begin{aligned} -\partial_t \Psi(t, \nu) = & \int_{\mathbb{R}^d} b(z, \nu) \cdot \partial_\mu \Psi(t, \nu)[z] - \frac{1}{2} |\sigma^\top(z, \nu) \partial_\mu \Psi(t, \nu)[z]|^2 \\ & + \frac{1}{2} \sigma \sigma^\top(z, \nu) : \partial_z \partial_\mu \Psi(t, \nu)[z] \nu(dz) \\ \Psi(T, \nu) = & g(\nu) \end{aligned}$$

- Similar to the master equation in mean-field games (J.F. Chassagneux et al., 2022; R. Carmona and F. Delarue, 2014)

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- Subsolutions are sufficient for log-efficiency.  
Uniqueness/regularity is active area (C. Wu and J. Zhang, 2020;  
A. Cosso et al., 2021; H.M. Soner and Q. Yan, 2022).  
Zero-viscosity limit studied in (M. Germain et al., 2021).

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Zero-viscosity limit studied in (M. Germain et al., 2021).
- Related equation used to prove quantitative rates of propagation of chaos in (P. E. Chaudru de Raynal and N. Frikha, 2021; J.F. Chassagneux et al., 2022)

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Zero-viscosity limit studied in (M. Germain et al., 2021).
- Related equation used to prove quantitative rates of propagation of chaos in (P. E. Chaudru de Raynal and N. Frikha, 2021; J.F. Chassagneux et al., 2022)
- Numerical methods for solving for the optimal control (J.P. Fouque and Z. Zhang, 2020; M. Laurière, 2021)

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# Numerical Example 1

$$dX_t^{i,N,y} = [-X_t^{i,N,y} + \frac{2}{N} \sum_{j=1}^N X_t^{j,N,y}] dt + \sigma dW_t^i, X_0^{i,N,y} = y$$

$$g(\nu) = \int_{\mathbb{R}} z^2 \nu(dz), \quad \text{target} = \mathbb{E} \left[ \exp \left( - \sum_{i=1}^N (X_T^{i,N,y})^2 \right) \right]$$

$$\begin{aligned} \nu(t, \nu, x) &= -2\sigma \left[ \frac{e^{2t}}{e^{2T}(1+\sigma^2) - e^{2t}\sigma^2} x \right. \\ &\quad \left. + \left( \frac{e^{2T}}{e^{2T}\sigma^2 - e^{2t}(\sigma^2 - 1)} - \frac{e^{2t}}{e^{2T}(1+\sigma^2) - e^{2t}\sigma^2} \right) \int z \nu(dz) \right] \end{aligned}$$

Example ( $\sigma = .5$ ,  $T = 1$ ,  $y = .2$ )

$N$	IS Est.	IS Rel.Err.	MC Est.	MC Rel.Err.	Exact
5	$2.3816 \cdot 10^{-1}$	$3.4601 \cdot 10^{-2}$	$2.3807 \cdot 10^{-1}$	1.0380	$2.3747 \cdot 10^{-1}$
10	$8.2550 \cdot 10^{-2}$	$2.7101 \cdot 10^{-2}$	$8.2551 \cdot 10^{-2}$	1.5721	$8.2412 \cdot 10^{-2}$
15	$2.8641 \cdot 10^{-2}$	$2.4001 \cdot 10^{-2}$	$2.8661 \cdot 10^{-2}$	2.1957	$2.8600 \cdot 10^{-2}$
20	$9.9373 \cdot 10^{-3}$	$2.2369 \cdot 10^{-2}$	$9.9165 \cdot 10^{-3}$	2.9589	$9.9254 \cdot 10^{-3}$
25	$3.4486 \cdot 10^{-3}$	$2.1310 \cdot 10^{-2}$	$3.4824 \cdot 10^{-3}$	3.9086	$3.4445 \cdot 10^{-3}$
30	$1.1968 \cdot 10^{-3}$	$2.0550 \cdot 10^{-2}$	$1.1951 \cdot 10^{-1}$	5.1415	$1.1954 \cdot 10^{-3}$
50	$1.7361 \cdot 10^{-5}$	$1.8934 \cdot 10^{-2}$	$1.7569 \cdot 10^{-5}$	14.5202	$1.7339 \cdot 10^{-5}$
80	$3.0339 \cdot 10^{-8}$	$1.8003 \cdot 10^{-2}$	$3.2426 \cdot 10^{-8}$	74.1871	$3.0289 \cdot 10^{-8}$

# Numerical Example 1

$$dX_t^{i,N,y} = [-X_t^{i,N,y} + \frac{2}{N} \sum_{j=1}^N X_t^{j,N,y}]dt + \sigma dW_t^i, X_0^{i,N,y} = y$$

$$g(\nu) = \int_{\mathbb{R}} z^2 \nu(dz), \quad \text{target} = \mathbb{E} \left[ \exp \left( - \sum_{i=1}^N (X_T^{i,N,y})^2 \right) \right]$$

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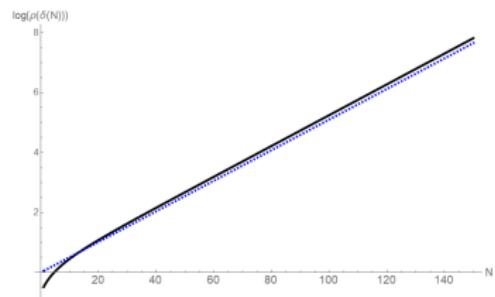
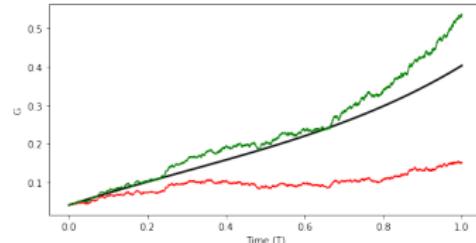


Figure: Left: A realization of  $G(\mu_T^N)$  (green) and  $G(\hat{\mu}_T^N)$  (red) against the second moment of the limiting McKean-Vlasov Equation (black) for  $T \in [0, 1]$  and  $N = 50$ . Right: The log-scale exact standard MC relative error as  $N$  increases (black).

# Numerical Example 2

$$dX_t^{i,N,y} = [-X_t^{i,N,y} + \frac{2}{N} \sum_{j=1}^N X_t^{j,N,y}]dt + \sigma dW_t^i, X_0^{i,N,y} = y$$

$$g(\nu) = \left| \int_{\mathbb{R}} z \nu(dz) \right|, \quad \text{target} = \mathbb{E} \left[ \exp \left( - \left| \sum_{i=1}^N X_T^{i,N,y} \right| \right) \right]$$

$$\nu(t, \nu, x) = -\sigma e^{T-t} \text{sign} \left( \int z \nu(dz) \right)$$

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Example ( $\sigma = .5$ ,  $T = 1$ ,  $y = .4$ )

$N$	IS Est.	IS Rel.Err.	MC Est.	MC Rel.Err.
5	$2.6644 \cdot 10^{-2}$	$3.7070 \cdot 10^{-1}$	$2.6653 \cdot 10^{-2}$	2.9930
10	$9.0804 \cdot 10^{-4}$	$3.1427 \cdot 10^{-1}$	$9.2009 \cdot 10^{-4}$	12.8032
15	$3.0269 \cdot 10^{-5}$	$2.6350 \cdot 10^{-1}$	$3.0645 \cdot 10^{-5}$	53.9410
20	$9.9533 \cdot 10^{-7}$	$2.2227 \cdot 10^{-1}$	$1.0625 \cdot 10^{-6}$	337.0401
25	$3.2471 \cdot 10^{-8}$	$1.8900 \cdot 10^{-1}$	$3.5791 \cdot 10^{-8}$	561.9164
30	$1.0537 \cdot 10^{-9}$	$1.6154 \cdot 10^{-1}$	$1.2253 \cdot 10^{-9}$	1016.1656

Importance Sampling vs. Standard Monte Carlo Estimate and Relative Error

# Numerical Example 3

$$dX_t^{i,N,y} = [-X_t^{i,N,y} + \frac{2}{N} \sum_{j=1}^N X_t^{j,N,y}]dt + \sigma dW_t^i, X_0^{i,N,y} = y$$

$$g(\nu) = \int_{\mathbb{R}} |z| \nu(dz), \quad \text{target} = \mathbb{E} \left[ \exp \left( - \sum_{i=1}^N |X_T^{i,N,y}| \right) \right]$$

$$\nu(t, \nu, x) = -\sigma \left[ e^{t-T} \text{sign}(x) + \left( e^{T-t} - e^{t-T} \right) \int \text{sign}(z) \nu(dz) \right]$$

Example ( $\sigma = .5$ ,  $T = 1$ ,  $y = .4$ )

$N$	IS Est.	IS Rel.Err.	MC Est.	MC Rel.Err.
5	$2.1355 \cdot 10^{-2}$	$4.7870 \cdot 10^{-1}$	$2.1382 \cdot 10^{-2}$	2.3263
10	$5.6703 \cdot 10^{-4}$	$6.0605 \cdot 10^{-1}$	$5.6806 \cdot 10^{-4}$	6.6374
15	$1.5087 \cdot 10^{-5}$	$7.2452 \cdot 10^{-1}$	$1.5040 \cdot 10^{-5}$	17.3337
20	$4.0157 \cdot 10^{-7}$	$8.3416 \cdot 10^{-1}$	$4.1260 \cdot 10^{-7}$	51.1058
25	$1.0687 \cdot 10^{-8}$	$9.4327 \cdot 10^{-1}$	$1.0296 \cdot 10^{-8}$	97.8820
30	$2.8470 \cdot 10^{-10}$	1.0524	$2.8800 \cdot 10^{-10}$	280.5945

Importance Sampling vs. Standard Monte Carlo Estimate and Relative Error

# Extensions and Future Work

## Future directions:

- Solve for the optimal controls on the fly
- Extend analysis to estimate probabilities
- Importance sampling scheme for long-time/metastable behavior (P. Dupuis et al., 2015)
- Moderate deviations based importance sampling (Z. B. and K. Spiliopoulos, 2022a; I. Gasteratos et al., 2022; F. Delarue and A. Tse, 2021)
- Joint small noise and large  $N$  limit (A. Budhiraja and M. Conroy, 2022)

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Thank you to my advisor, Kostas Spiliopoulos, to my coauthor and friend Max Heldman, and to all of you for listening!

## Questions?

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## Definition

Let  $\{X^N\}_{N \in \mathbb{N}}$  be a sequence of random variables taking values in a complete separable metric space  $\mathcal{X}$ . We say that  $\{X^N\}$  satisfies the *large deviation principle* with rate function  $I$  if there exists  $I : \mathcal{X} \rightarrow [0, \infty]$  with compact level sets such that

- 1  $\limsup_{N \rightarrow \infty} \frac{1}{N} \log \mathbb{P}(X^N \in F) \leq -\inf_{x \in F} I(x), \forall F \subset \mathcal{X}$   
such that  $F$  is closed
- 2  $\liminf_{N \rightarrow \infty} \frac{1}{N} \log \mathbb{P}(X^N \in G) \geq -\inf_{x \in G} I(x), \forall G \subset \mathcal{X}$   
such that  $G$  is open.

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such that  $G$  is open.

- Idea: For a Borel set  $E \subset \mathcal{X}$ , we have

$$\mathbb{P}(X^N \in E) \sim \exp\left(-N \inf_{x \in E} I(x)\right) \text{ as } N \rightarrow \infty.$$

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- 2  $\liminf_{N \rightarrow \infty} \frac{1}{N} \log \mathbb{P}(X^N \in G) \geq -\inf_{x \in G} I(x), \forall G \subset \mathcal{X}$   
such that  $G$  is open.

- Idea: For a Borel set  $E \subset \mathcal{X}$ , we have

$$\mathbb{P}(X^N \in E) \sim \exp\left(-N \inf_{x \in E} I(x)\right) \text{ as } N \rightarrow \infty.$$

- So if  $X^N \xrightarrow{d} X$ , and  $X \notin E$ ,  $\inf_{x \in E} I(x) > 0$  characterizes the exponential rate of decay of the rare event  $X^N \in E$ .

# Dawson-Gärtner LDP (1)

$$dX_t^{i,N} = b(X_t^{i,N}, \mu_t^N)dt + \sigma(X_t^{i,N})dW_t^i, \quad \mu_t^N := \frac{1}{N} \sum_{i=1}^N \delta_{X_t^{i,N}}$$

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Theorem (D. A. Dawson and J. Gärtner, 1987)

$\{\mu^N\}_{N \in \mathbb{N}}$  satisfies the LDP with rate function  
 $I^{DG} : C([0, T]; \mathcal{P}(\mathbb{R}^d)) \rightarrow [0, +\infty]$  given by

$$I^{DG}(\mu) = \frac{1}{2} \int_0^T \sup_{\phi \in C_c^\infty(\mathbb{R}^d) : \text{den} \neq 0} \frac{|\langle \dot{\mu}(t) - L_{\mu(t)}^* \mu(t), \phi \rangle|^2}{\langle \mu(t), \|\nabla^\top \phi \sigma \sigma^\top \nabla \phi\|^2 \rangle} dt$$

$$L_\nu \phi(x) := b(x, \nu) \cdot \nabla \phi(x) + \frac{1}{2} \sigma \sigma^\top(x) : \nabla^2 \phi(x), \quad \nu \in \mathcal{P}(\mathbb{R}^d)$$

if  $\phi \mapsto \langle \mu, \phi \rangle$  is absolutely continuous in the sense of distributions  
and  $\mu(0) = \lim_{N \rightarrow \infty} \mu_0^N$ , and  $I^{DG}(\mu) = +\infty$  otherwise.

# Dawson-Gärtner LDP (2)

$$dX_t^{i,N} = b(X_t^{i,N}, \mu_t^N)dt + \sigma(X_t^{i,N})dW_t^i, \quad \mu_t^N := \frac{1}{N} \sum_{i=1}^N \delta_{X_t^{i,N}}$$

$$I^{DG}(\mu) = \frac{1}{2} \int_0^T \sup_{\phi \in C_c^\infty(\mathbb{R}^d): \text{den} \neq 0} \frac{|\langle \dot{\mu}(t) - L_{\mu(t)}^* \mu(t), \phi \rangle|^2}{\langle \mu(t), \|\nabla^\top \phi \sigma \sigma^\top \nabla \phi\|^2 \rangle} dt$$

- Proof method: General theorems on projective limits for large deviation systems and LDPs on dual vector spaces to get the result for IID diffusion processes, then Girsanov's theorem to show closeness to an IID system

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- Can be expressed in terms of derivatives of the free energy associated to the limiting McKean-Vlasov Equation viewed as a gradient flow on Wasserstein space in some settings (S. Adams et al., 2013). In gradient setting, connected to HJB Equations in (J. Feng and G. Kurtz, 2006)

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- Transition events for Dawson's system estimated in (J. Garnier)

# Buhiradja-Dupuis-Fischer LDP (1)

$$dX_t^{i,N} = b(X_t^{i,N}, \mu_t^N)dt + \sigma(X_t^{i,N}, \mu_t^N)dW_t^i, \quad \mu_t^N := \frac{1}{N} \sum_{i=1}^N \delta_{X_t^{i,N}}$$

Theorem (A. Budhiraja et al., 2012)

$\{\mu^N\}_{N \in \mathbb{N}}$  satisfies the LDP with rate function  
 $I : C([0, T]; \mathcal{P}(\mathbb{R}^d)) \rightarrow [0, +\infty]$  given by

$$I(\mu) = \inf_{u \in \mathcal{U}: \mu(t) = \mathcal{L}(\hat{X}_t^u), t \in [0, T]} \frac{1}{2} \mathbb{E} \left[ \int_0^T |u(t)|^2 dt \right]$$

$$\begin{aligned} d\hat{X}_t^u &= [b(\hat{X}_t^u, \mathcal{L}(\hat{X}_t^u)) + \sigma(\hat{X}_t^u, \mathcal{L}(\hat{X}_t^u))u(t)]dt \\ &\quad + \sigma(\hat{X}_t^u, \mathcal{L}(\hat{X}_t^u))dW_t, \quad \hat{X}_0^u \sim \nu \end{aligned}$$

if  $\mu(0) = \nu = \lim_{N \rightarrow \infty} \mu_0^N$  and  $I(\mu) = +\infty$  otherwise. Here  $\mathcal{U}$  is an appropriate class of stochastic controls  $u : [0, T] \rightarrow \mathbb{R}^m$

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## Buhiradja-Dupuis-Fischer LDP (2)

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- $I = I^{DG}$  proved even when  $\sigma$  depends on  $\mu$  in (Z. B. and K. Spiliopoulos, 2023)

# Lions Differentiation

Consider:

$$\mathcal{P}_2(\mathbb{R}^d) := \{\mu \in \mathcal{P}(\mathbb{R}^d) : \int_{\mathbb{R}^d} |x|^2 \mu(dx) < \infty\}.$$

## Lions Derivative (R. Carmona and F. Delarue, 2018) Section 5.2

- We say  $u : \mathcal{P}_2(\mathbb{R}^d) \rightarrow \mathbb{R}$  is **Lions-differentiable** at  $\mu_0 \in \mathcal{P}_2(\mathbb{R}^d)$  if there exists  $\tilde{u} : L^2(\tilde{\Omega}, \tilde{\mathcal{F}}, \tilde{\mathbb{P}}; \mathbb{R}^d) \rightarrow \mathbb{R}$  such that  $\tilde{u}(X) = u(\mathcal{L}(X))$ ,  $\forall X \in L^2(\tilde{\Omega}, \tilde{\mathcal{F}}, \tilde{\mathbb{P}}; \mathbb{R}^d)$  and  $\tilde{u}$  is Fréchet differentiable at  $X_0$  with  $\mu_0 = \mathcal{L}(X_0)$ .

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- There is a  $\mu_0$  a.s. unique deterministic measurable function  $\xi : \mathbb{R}^d \rightarrow \mathbb{R}^d$  such that  $D\tilde{u}(X_0) = \xi(X_0)$ . We denote this equivalence class of  $\xi \in L^2(\mathbb{R}^d, \mu_0; \mathbb{R}^d)$  by  $\partial_\mu u(\mu_0)$  and call  $\partial_\mu u(\mu_0)[\cdot] : \mathbb{R}^d \rightarrow \mathbb{R}^d$  the **Lions derivative** of  $u$  at  $\mu_0$ .

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- Note that this definition is independent of the choice of  $X_0$ ,  $\tilde{u}$ , and  $(\tilde{\Omega}, \tilde{\mathcal{F}}, \tilde{\mathbb{P}})$ .

# Linear Functional Derivative

## Definition: Linear Functional Derivative

Let  $p : \mathcal{P}_2(\mathbb{R}) \rightarrow \mathbb{R}$ ,  $\mathcal{P}_2(\mathbb{R}) = \{\mu \in \mathcal{P}(\mathbb{R}) : \int |z|^2 \mu(dz) < \infty\}$ .

$(z, \mu) \ni \mathbb{R} \times \mathcal{P}_2(\mathbb{R}) \mapsto \frac{\delta}{\delta m} p(\mu)[z] \in \mathbb{R}$  is called the **Linear Functional Derivative** of  $p$  if for all  $\nu_1, \nu_2 \in \mathcal{P}_2(\mathbb{R}^d)$ :

$$p(\nu_2) - p(\nu_1) = \int_0^1 \int_{\mathbb{R}} \frac{\delta}{\delta m} p((1-r)\nu_1 + r\nu_2)[z](\nu_2(dz) - \nu_1(dz))$$

Example:  $b(\mu) = \int \beta(z) \mu(dz)$ , then  $\frac{\delta}{\delta m} b(\mu)[z] = \beta(z)$ .

Relationship:  $\partial_\mu p(\mu) = \nabla_z \frac{\delta}{\delta m} p(\mu)[z]$ .

# Lions Differentiation: Empirical Projection

Proposition- 5.35 and 5.91 of (R. Carmona and F. Delarue, 2018)

For  $g : \mathcal{P}_2(\mathbb{R}^d) \rightarrow \mathbb{R}$  twice continuously Lions differentiable with  $C^1$  first Lions derivative, we can define the empirical projection of  $g$ , as  $g^N : \mathbb{R}^{dN} \rightarrow \mathbb{R}$  given by

$$g^N(x_1, \dots, x_N) := g\left(\frac{1}{N} \sum_{i=1}^N \delta_{x_i}\right) = g(\mu_x^N).$$

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Then  $g^N$  is twice differentiable on  $(\mathbb{R}^d)^N$ , and for each  $x_1, \dots, x_N \in \mathbb{R}^d$ ,  $(i, j) \in \{1, \dots, N\}^2$ :

$$\nabla_{x_i} g^N(x_1, \dots, x_N) = \frac{1}{N} \partial_\mu g(\mu_x^N)[x_i]$$

and

$$\nabla_{x_i} \nabla_{x_j} g^N(x_1, \dots, x_N) = \frac{1}{N} \partial_z \partial_\mu g(\mu_x^N)[x_i] \mathbb{1}_{i=j} + \frac{1}{N^2} \partial_\mu^2 g(\mu_x^N)[x_i, x_j].$$